

Do Campaigns Drive Partisan Turnout?

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Abstract Although campaign strategy often, and perhaps increasingly, emphasizes the mobilization of core supporters, we know little about whether campaigns affect the partisan complexion of the electorate. We examine whether the balance of Democratic and Republican voters depends on the balance of campaign activity, the popularity of the incumbent president, and the state of the economy. Drawing on time-series cross-sectional data from state exit polls, we demonstrate that the partisan composition of voters depends on campaign activity more than on the political and economic fundamentals.

Keywords Voter turnout · Political participation · Parties · Elections · Campaigns · United States

The question of partisan mobilization—did more Democrats or Republicans vote in the last election?—is central to election outcomes. The lore of campaigns has long divided the electorate into “base” voters, who are reliable supporters but must be mobilized to vote, and “swing” voters, who are likely to be at the polls but must be persuaded to support the campaign. Mobilizing the base has always been part of campaigns, but recently it has figured more prominently in campaign strategy. Campaigns have assembled large databases of information about voters, better enabling them to target voters who are predisposed to support them and get those voters to the polls. Many commentators attributed Al Gore’s surprisingly strong

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performance in 2000 to an eleventh-hour mobilization of union supporters, and likewise attributed Republican successes in 2002 and 2004 to the Republican get-out-the-vote operation, or “72 Hour Plan.”

Despite this interest in partisan mobilization, our understanding of the effects that campaigns have on turnout is still in its infancy. Is mobilization an important consequence of campaigns, or do campaigns exert their influence, if at all, through the persuasion of swing voters? Are these effects independent of the “fundamentals”—such as presidential approval and the state of the economy—that are known to influence election outcomes? Is there even significant variation across elections in the proportions of Democratic and Republican voters?

Answers to these questions speak to debates about the scope and magnitude of campaign effects. While it is clear that some kinds of campaign activities can get out the vote, it is less clear which voters “get out.” We do not know very much about how campaign activities affect change at the level of the electorate itself. The answers also have implications for the substance of politics, because campaigns that pursue a mobilization strategy may have to take more ideologically extreme positions to appeal to the party faithful.

We isolate the aggregate effect of campaign activity on “partisan turnout”—that is, the proportion of Democrats and Republicans who actually turn out to vote—from other variables, such as the fundamentals, that might also be bringing voters of one party or the other to the polls. We demonstrate that campaigns do, in fact, affect the turnout of their core supporters, above and beyond any effects of the fundamentals.

Political Mobilization and Partisan Turnout

Much research has investigated whether turnout depends on factors specific to an election year, in addition to its well-known relationship to individual-level attributes (Wolfinger and Rosenstone 1980). Simple excitement about the campaign is one such factor: races higher on the ballot (such as president or governor) can be a draw for voters (Arcelus and Meltzer 1975; Campbell 1960; Gilliam 1985), as can competitive races further down the ballot (Cox and Munger 1989). Campaigns, parties, and interest groups can also mobilize voters to vote (Ansolabehere and Iyengar 1995; Caldeira and Patterson 1982; Cox and Munger 1989; Jackson 2002; Rosenstone and Hansen 1993). Although some of the relevant evidence is ambiguous (Paul 2002), and the effects of mobilization are often difficult to distinguish from simple excitement, the best efforts to separate the two suggest that certain mobilization techniques increase turnout (Green and Gerber 2004).

Our focus is not on the percentage of people who voted, but rather the partisan complexion of those voters. We are interested in what proportion of voters identifies as Democratic or Republican, and how that proportion changes in response to short-term forces. Despite the received wisdom that campaigns mobilize partisans, systematic studies of partisan turnout have been rare. The most relevant study is Holbrook and McClurg’s (2005), which identifies the “mobilization of core supporters” as a central contribution of the 1992, 1996, and 2000 presidential

campaigns. In particular, spending by the parties themselves—specifically transfers by the national parties to the states—affects partisan turnout. Shaw (1999) also finds evidence of mobilization in his study of the 1988–1996 presidential campaigns, but his measures are indirect; he interprets campaign activity as mobilizing voters when its effect on one party’s vote share increases as a state’s partisan balance tilts more in favor of that party. Finally, Campbell (1997) has argued that partisan turnout helps explain why presidents almost always lose seats and votes in midterm elections: a “surge” in loyal partisans in presidential years is followed by a “decline” of those partisans in midterm elections, a phenomenon exacerbated by current events and presidential popularity. But his analysis sets aside the question of campaign influence on turnout, at least beyond the mobilizing effects of the presidential race writ large.

Thus, evidence for short-term shifts in partisan turnout is tentative, even as turnout effects loom large in political science theories and popular accounts of elections. Our approach helps fill this gap in two key ways. First, we expand the scope of previous analyses to include both presidential and midterm elections from 1988 to 2006. The inclusion of midterm elections allows us to consider a broader range of campaign activity—including contests for governor and Congress independent of a presidential race—which helps us gauge the impact of lower-salience campaigns. Second, we explicitly consider the impact of “fundamentals” such as the state of the economy and the approval ratings for prominent political leaders. These fundamentals might actually produce any apparent campaign effects, since a promising political environment could make it easier to recruit quality candidates, raise money, and mount campaigns that will attract sympathetic partisans to the polls. Indeed, considering the modest effects of campaigns described by existing literature, we cannot ignore the possibility that mobilization efforts simply “piggy-back” on the broader political climate. By expanding the theoretical perspective and empirical analysis, we can better describe partisan turnout and better identify its origins.

What Affects Partisan Turnout?

Turnout is conditional for many voters, who shift in and out of the electorate from election to election (Sigelman and Jewell 1986). Because these voters comprise only a portion of the total electorate, shifts in partisan composition will be evident but not necessarily large. For partisan composition to shift, voters of one party must increase relative to voters of the other party. What might create asymmetries in the enthusiasm or motivation of partisans?

The explanation that has received the most scholarly attention is mobilization by political campaigns. Candidates and parties have an incentive to spend their money efficiently, and mobilization of partisans may prove more efficient than other ways of winning votes, such as the persuasion and mobilization of non-partisans (see Holbrook and McClurg 2005). Campaign mobilization should increase participation for two reasons. First, it provides loyal partisans who are “marginal” voters—that is, whose participation is not assured but is at least possible—with a variety of

useful resources: information about the candidates, assistance in registering to vote and getting to the polls, and reminders to turn out on Election Day itself. Thus, mobilization helps to reduce the well-documented costs of voting (Verba et al. 1995). Mobilization also gives voters a reason to care about the outcome, usually by contrasting their party's candidate with the opposition. This can increase general excitement about the race, thereby boosting the psychological benefits of voting. Thus, asymmetries in the parties' spending may lead to differences in partisans' perceived costs and benefits, both real and psychological. The party that spends more will increase its share of the electorate.

A second short-term effect on partisan turnout is an "activation" that derives from the "nature of the times," to borrow a phrase from Converse (1964). The mechanism here again involves the psychological benefits of voting. Some partisans are similar to sports fans, who are more likely to attend games when their team is winning (e.g., Coates and Humphreys 2005). When things are going well for their party, these partisans will be more enthused, and enthusiasm foments participation (Marcus et al. 2000). By contrast, when things are going poorly, they will become fair-weather fans, retaining their party identification but failing to act on it.¹ This effect should be independent of the quality of the party's candidates or the strength of its campaigns—conceptually, at least, it derives only from the political context in that election year. Fans will watch a winning team even if they dislike some of its players.

How do we know if a party's fortunes are waxing or waning? We focus our attention on two indicators: approval of a party's incumbent office-holders, and the health of the economy. Each of these factors figures prominently in the literature on forecasting elections (e.g., Lewis-Beck and Tom Rice 1992), suggesting that they serve as focal points for voters. We expect that when a party's leaders, such as Senators or the president, are well-liked, the partisan composition of the electorate will tilt in favor of that party. The effect of the economy should also depend on the party of the incumbent. When the economy is healthy, the partisan composition of the electorate should shift in favor of the incumbent's party. An unhealthy economy will similarly benefit the out-party.

Both mobilization and activation have similar effects on the turnout calculus of individual voters; for instance, both influence the purely psychological sense of enthusiasm. In fact, the difference between mobilization and activation is less about voter psychology than about elite control. Mobilization is far more open to short-term manipulation by political actors, while the fundamentals that drive activation are almost entirely outside the control of any individual campaign. Indeed, some of the most prominent theories of activation downplay the role of campaigns, at least beyond the presidential contest (Campbell 1997, 1960). By contrast, campaigns have much more control over their mobilization efforts—where resources are spent, how they are spent, who is targeted, etc. Of course, activation and mobilization are potentially entwined because a popular party may have strong campaigns as well

¹ Of course, many people maintain loyalty to a group even when it is chronically unsuccessful. This is why even Wrigley Field can sell out. Diehard fans, like habitual voters, are likely to participate even when their team's prospects are dim.

(Jacobson and Kernell 1983). This means that the apparent effect of campaigns on partisan turnout might instead be the product of forces beyond the campaign's control. Addressing this issue requires looking at both activation and mobilization in the same analysis.

Ultimately, it matters to our understanding of campaigns and campaign influence whether mobilization affects partisan turnout and how large that effect is likely to be. If campaigns can mobilize the party faithful, then it makes sense to cater to these voters, which could entail a more ideological party platform, among other things. If campaigns do not affect partisan turnout, then it makes sense for parties to cater to swing voters and moderate their platform. By assessing the impact of mobilization alongside the impact of political and economic context, we can develop a more complete picture of campaign influence.

Partisan Turnout in the States, 1988–2006

To evaluate these expectations, we measure change in the partisan composition of state electorates over time, and demonstrate how this change is related to incumbent approval, economic health, and campaign activity. States differ across all of these dimensions, not only cross-sectionally but temporally, thereby providing us more analytical leverage than a single time-series at the national level or within a single state could provide. This is particularly true with regard to campaign activity, which varies widely among states because of differences in the competitiveness of races. Aggregating campaign activity to the national level would wash out this variation by obscuring the differences in campaign activity between, for example, battleground and “blackout” states in presidential campaigns (see Gimpel et al. 2007; Shaw 2006).

We measure partisan composition with state-level exit polls. State-level polls were first conducted in 1988 and have been conducted in almost every election year since. The exception is 2002, when the exit poll was exclusively a national sample. In midterm election years, exit polls were typically conducted only in a subset of states, ranging from 24 states in 1988 to 41 in 1998. In 2002, the national sample contained respondents from 42 states, but in many cases there were not enough respondents for reliable state-level estimates. As a starting point for the analysis, we included only seven states in which the sample numbered at least 600; this was the approximate minimum sample size in the 2000 and 2004 state-level exit polls.² This leaves us with a total sample size of 359 state-years. Appendix lists the states included in each year.

² We omitted New York, which, despite its large sample size ($N = 1088$), was clearly an outlier in terms of its partisan turnout. In particular, its partisan turnout in 2002 was skewed well toward the Republican Party. This likely stems from the exit poll's sampling within New York. Although precise information about which precincts were included is not available in the publicly released data, no exit poll respondents are designated as living in cities with more than 500,000 people, suggesting that no precincts from New York City were included in the sample. This could help account for the skew in the sample's partisan composition.

Within each state's poll, we calculated the proportion of voters who identified as Democratic, Republican, or Independent, using the sample weights in generating these proportions.³ Our primary measure of partisan composition is the relative balance between Democrats and Republicans, operationalized as the percent Democratic minus the percent Republican (see also Holbrook and McClurg 2005). Figure 1 presents that figure for each of the 50 states in each available election year.

An important feature of Fig. 1 is the sheer diversity among states, in terms of not only the average Democratic or Republican advantage (which was expected) but also the trends in this advantage. Some states, such as Florida, are quite evenly balanced and change little over time. Other states manifest clear trends in favor of one party, such as Alabama and Louisiana for the Republicans or New Jersey and Illinois for the Democrats. In other cases, such as Arkansas, the trend is not completely monotonic. Thus, although party identification is itself a fairly stable attribute (Green et al. 2002), the partisan composition of the electorate varies over time in many states. The next task is to explain this variation.

To do so, we operationalize the three concepts discussed previously, beginning with campaign activity. Because our estimate of the partisan composition is at the state level, we need to capture as best as possible the various kinds of campaign activity in that state, both in midterm and in presidential years. To do so, we computed the total general election spending of all House candidates as well as that of gubernatorial and senatorial candidates when applicable.⁴ In addition, we included the total television advertising of the presidential candidates (in gross rating points) in each state, as well as the total number of visits that the presidential candidates made to each state (see Shaw 1999, 2006). Finally, for presidential years from 1992 through 2004, we also included spending by the state parties on voter mobilization.⁵ To measure the balance of campaign activity between the parties, we first standardized each of these measures so that it ranged from -1 (complete Republican advantage) to $+1$ (complete Democratic advantage).⁶ We then averaged these standardized measures to make an overall index of the partisan balance of

³ The exit polls do not allow independents to designate a preferred party and thus do not capture independent "leaners." We explored the impact of incorporating leaners into our measure of partisan turnout using two datasets: validated turnout in the American National Election Study (ANES) (available for 1964, 1976, 1978, 1980, 1984, 1986, 1988, and 1990), and validated turnout in the 2006 Cooperative Congressional Election Study (CCES). The correlation between our measure and the same measure with leaners was $r = 0.98$ across years in the ANES and $r = 0.94$ across states in the CCES (using only states with at least 200 cases to reduce sampling noise).

⁴ House spending data from 1988 through 1998 were graciously provided by Gary Jacobson of the University of California at San Diego. The remaining House spending data and all the Senate spending data come from the Federal Election Commission. Gubernatorial spending data from 1988 through 2004 come from the Gubernatorial Campaign Expenditures database (Beyle and Jensen 2007); the data for 2006 come from the websites of the various states.

⁵ Ray La Raja of the University of Massachusetts, Amherst graciously provided these data. They are calculated as dollars per 1,000 eligible voting persons. Since we do not have data for 1988, the campaign balance measure for that year does not include this measure.

⁶ For example, to compute the balance of campaign spending for Democrats (D) and Republicans (R), we calculated: $(D - R)/(D + R)$. This has the advantage of generating comparable measures even though the underlying indicators (spending, ads, visits, etc.) are on different metrics.

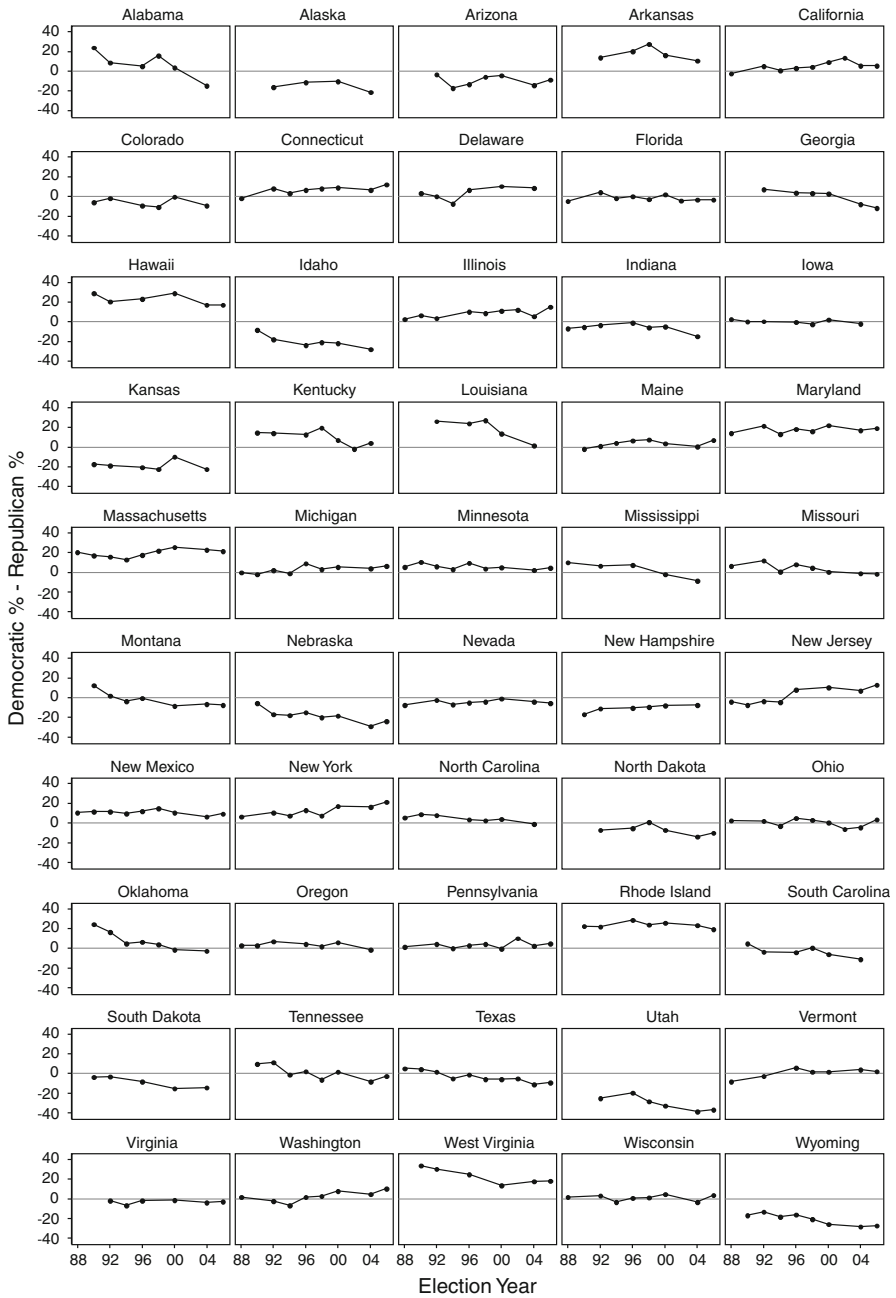


Fig. 1 Democratic advantage among exit poll voters, by state and year (1988–2006)

campaign activity, and multiplied the index by 100 to make it comparable to the measures described below. We refer to this measure as “campaign advantage.”⁷

To provide an initial depiction of this measure and its apparent effect, Fig. 2 displays the bivariate relationship between campaign advantage and the Democratic partisan advantage in each of the 50 states. The datapoints in each scatterplot correspond to election years. The least squares fit line and confidence intervals indicate whether the overall relationship is, as we expect, positive. In most cases (37 out of 50), the slope of the fit line is positive, suggesting that Democrats increase their share of voters as their campaigning increases relative to Republicans’. This confirms our expectation, although of course a more rigorous test is necessary.

The second concept is partisan advantage in incumbent approval. We drew on polling data from the U.S. Officials’ Job Approval Ratings Project (Niemi et al. 2007), extracting the percent who approved of the sitting governor, senators, and president in the state poll taken closest to the election. Our estimates of approval, including approval of the president, are therefore from state samples. Presidential approval is our primary measure. The president’s visibility as a party leader should make his approval rating more consequential for partisan turnout, although we will investigate approval of governors and senators as well. We multiplied the approval rating by -1 if the incumbent was a Republican. The resulting measure’s theoretical range is -100 (unanimous approval of a Republican incumbent) to $+100$ (unanimous approval of a Democratic incumbent). These approval measures capture the “nature of the times” by isolating each incumbent’s relative overall popularity in each state.

The chief limitation of the presidential approval measure is missing data. Because some states do not have presidential approval data specific to a particular election year, the resulting sample size is 270 state-years instead of the 360 for which we have exit poll data. Thus, we constructed a second measure of presidential approval that imputed approval for these missing cases. For each state, we computed the average deviation between that state’s presidential approval and presidential approval in the nation as a whole (in both cases, drawing on the polls closest to the election). Then we imputed state-level approval, when necessary, as the national approval minus the average difference between state and national polls. For the cases that did have valid state-level approval data, the correlation between the imputed measure and the actual measure is $r = 0.98$, suggesting that this imputation strategy creates plausible values for these missing cases. We employ both the original and imputed measures in our analysis.

The final concept is the health of the economy. To construct a state-specific measure, we calculated the percentage change in each state’s per capita disposable income (in 2006 dollars), comparing the election year to the previous year. Similar

⁷ Although this measure combines different types of campaign activities by different campaigns, some sort of combination is necessary in order to consider both presidential and midterm elections in the same analysis. Of course, partisan turnout might respond to specific types of campaign activities or activities by certain campaigns, rather than the sum of campaign activity in general. Thus, if anything, this measure imprecisely measures our construct of interest and understates the effect of campaigns.

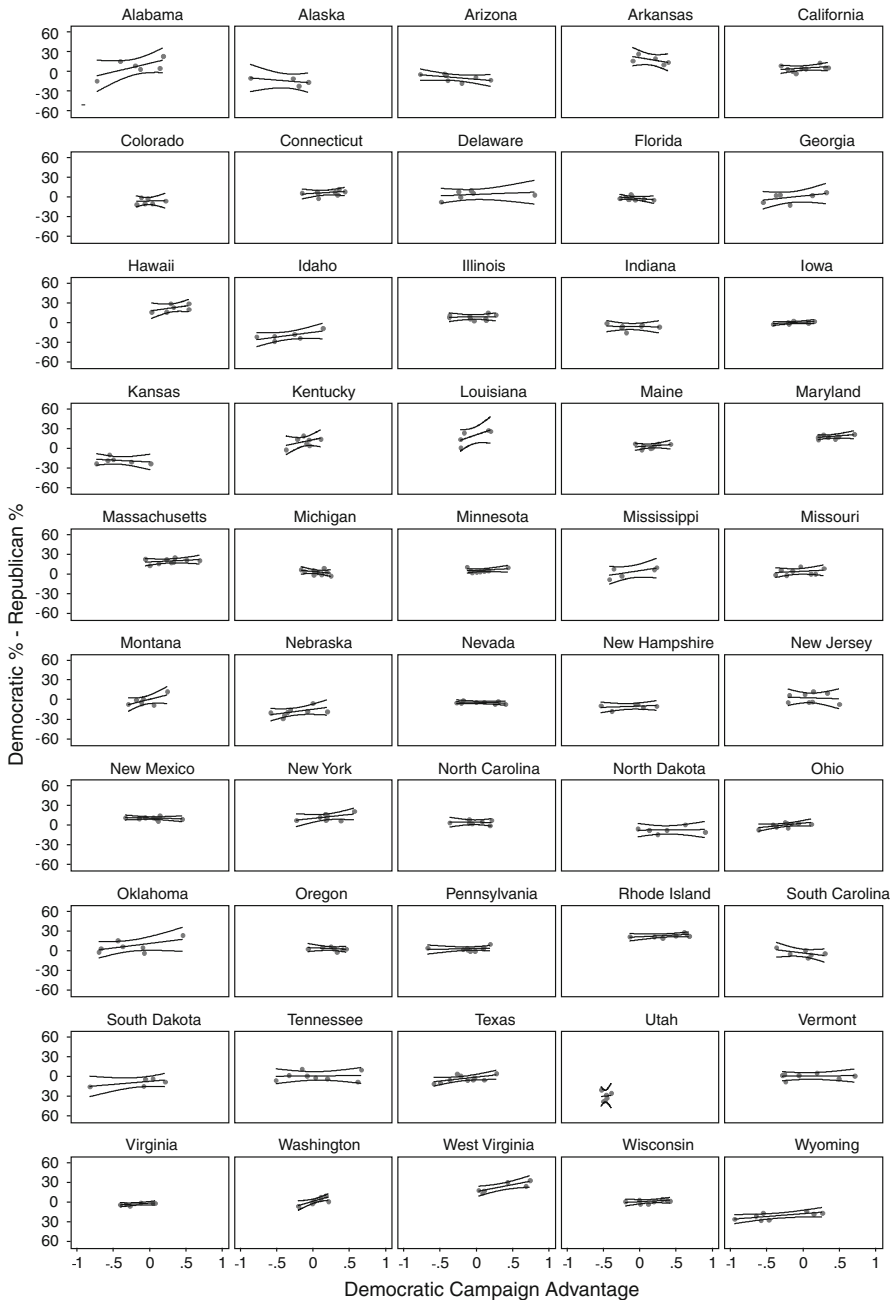


Fig. 2 Democratic advantage in partisan turnout and campaigning, by state

measures feature prominently in predictive models of election outcomes. This measure was also multiplied by -1 if the incumbent was a Republican. The key question is which incumbent office-holder is more likely to be held accountable for

Table 1 Time-series cross-sectional models of partisan turnout, 1988–2006

	With original approval			With imputed approval		
	1	2	3	4	5	6
Campaign activity	0.05** [0.02]	0.04* [0.02]	0.05** [0.02]	0.06** [0.01]	0.05** [0.01]	0.06** [0.01]
Presidential approval	0.06 [0.05]	0.07 [0.04]	0.06 [0.05]	0.05 [0.04]	0.06+ [0.03]	0.07 [0.04]
State economic change	0.01 [0.23]	0.04 [0.19]	−0.14 [0.24]	0.07 [0.15]	−0.01 [0.15]	−0.03 [0.19]
President's party	−2.24 [2.39]	−3.98* [1.91]	0.72 [3.36]	−1.95 [1.93]	−3.28* [1.53]	−3.00 [2.07]
State-level macropartisanship		0.12+ [0.06]			0.09+ [0.05]	
National economic change			0.93 [0.77]			−0.23 [0.18]
Constant	1.73** [0.06]	1.90** [0.45]	1.59** [0.12]	1.83** [0.13]	2.18** [0.29]	1.66** [0.17]
Observations	269	194	269	359	270	359
Number of states	50	48	50	50	48	50
R-squared	0.09	0.10	0.10	0.10	0.14	0.11

Cell entries are least squares coefficients, with robust standard errors in parentheses and fixed effects for states. The dependent variable is the difference between the percentage of Democrats and Republicans in the exit poll in each state-year, coded such that higher values indicate a Democratic advantage. ⁺ $p < 0.10$; * $p < 0.05$; ** $p < 0.01$

the performance of the state's economy. We constructed two versions, one based on the party of the incumbent governor and another based on the party of the incumbent president. Both measures range from roughly -20 to $+20$.

Using these measures, we estimated time-series cross-sectional models of the balance of Democratic and Republican partisans (i.e., the measure in Fig. 1), including each of these independent variables.⁸ Table 1 presents six different specifications: three with observed presidential approval and three with imputed approval for those cases where approval was missing. Every model includes robust standard errors and fixed-effects for states, and all are estimated in State 10.0.⁹ Table 1 displays the results.

Model 1 includes the variables described above and includes only those cases for which we have measured presidential approval. The results suggest that campaign activity does have a statistically significant effect on partisan turnout—Democratic

⁸ Our models also include a dummy variable for the party of the president, coded -1 (Republican) and $+1$ (Democrat), which accounts for the coding of the economy and presidential approval variables.

⁹ Models with panel-corrected standard errors and a first-order auto-regressive component generate results very similar to those presented here (Beck and Katz 1995; Wilson and Butler 2007). These results are available on request.

partisans are more prevalent, relative to Republican partisans, when Democrats have an advantage in campaigning. The same could not be said for either of the “fundamentals.” The effects of presidential approval and state economic change are both statistically insignificant, and the effect of economic change is also substantively small.

Model 2 tests whether the effect of campaign activity is an artifact of shifts in state partisanship that occur prior to the election. If campaigns spend more in those states where the underlying partisanship has shifted in their favor, then the correlation between campaign activity and turnout could be spurious. To test this idea, we included state-level macropartisanship, which is calculated in a similar fashion and measured prior to the election (Erikson et al. 1989).¹⁰ Since this variable is not available after 2002, the scope of the analysis shrinks somewhat. Its effect is positive and statistically significant at the 0.10 level: unsurprisingly, Democrats are more prevalent at the polls in states where they are simply more prevalent. Nonetheless, the effect of campaign activity remains essentially unchanged. Model 3 tests a different hypothesis: that partisan turnout responds to the state of the *national* economy, in “sociotropic” fashion, rather than the economy closer to home. This specification changes the magnitude of some coefficients, but leaves untouched both the effect of campaign activity and the substantive conclusions to be gleaned from the model. Finally, Models 4–6 employ imputed presidential approval for those cases where approval is missing. The results are similar to Models 1–3.¹¹

In sum, presidential approval and the state of the economy are not significantly associated with partisan turnout, while the effect of campaign activity is statistically significant and robust to alternative specifications. If we use the coefficient from Model 4—which makes best use of the data—the magnitude of this campaign effect suggests that a shift from a complete Republican campaign advantage (the sample minimum of –100) to a complete Democratic advantage (the sample maximum of +100) would increase the Democrats’ relative share of voters by 13 percentage points. Of course, such a large shift in activity is implausible. A more reasonable shift in campaigning—say, one standard deviation—would increase the Democrats’ share by approximately two points.

Do the effects we have found differ between midterm and presidential election years? There is a considerable literature on the surprisingly consistent seat loss for the president’s party in midterm elections, and this seat loss may derive in part from changes in the correlates of partisan turnout. Accordingly, we estimated Models 3 and 6 from Table 1, including a set of interactions between each covariate and presidential election year. Table 2 presents these results. The effects of campaign activity, presidential approval, and national economic change do not have distinguishable effects in presidential versus midterm years. The coefficients for

¹⁰ These data, calculated for 1988–2003, are available at <http://php.indiana.edu/~wright1/>.

¹¹ In alternative models, we replaced presidential approval with approval of the governor or either of the senators. None of these measures of approval had significant effects. At a reviewer’s suggestion, we also investigated whether there were interactions between campaign activity and either presidential approval or the state of the economy. We uncovered no statistically significant interactions.

Table 2 Models of partisan turnout, with interactions for presidential election years

	Original approval	Imputed approval
Campaign activity	0.04* [0.02]	0.07** [0.01]
Presidential approval	0.06 [0.07]	0.04 [0.06]
State economic change	-0.14 [0.33]	0.65* [0.31]
President's party	2.02 [5.63]	-4.51 [3.07]
National economic change	2.04 ⁺ [1.17]	-0.31 [0.27]
Campaign activity × presidential year	0.01 [0.02]	-0.01 [0.02]
Presidential approval × presidential year	-0.03 [0.09]	0.04 [0.06]
State economic change × presidential year	-0.06 [0.42]	-0.69 [0.43]
President's party × presidential year	7.58 [5.04]	1.71 [3.18]
National economic change × presidential year	1.15 [1.49]	0.25 [0.31]
Presidential year	1.92* [0.91]	0.44 [0.62]
Constant	0.27 [0.66]	1.27** [0.45]
Observations	269	359
Number of states	50	50
R-squared	0.19	0.18

Cell entries are least squares coefficients, with robust standard errors in parentheses and fixed effects for states. The dependent variable is the difference between the percentage of Democrats and Republicans in the exit poll in each state-year, coded such that higher values indicate a Democratic advantage.

⁺ $p < 0.10$; * $p < 0.05$; ** $p < 0.01$

the respective interaction terms are small in magnitude and statistically insignificant. There is some evidence that the effects of state economic change are larger in midterm years—where the effect is positive, as hypothesized, and statistically significant—than in presidential years, where the effect appears to be smaller. This finding emerges mainly in the model with imputed approval, although the interaction term is not quite statistically significant at conventional levels ($p = 0.12$). Nevertheless, our central finding—that campaign activity is associated with partisan turnout—is consistent across both types of election years.

Thus far, we have combined Democratic and Republican campaigning, as well as the proportion of Democratic and Republican voters, into single measures. As such, we cannot discern any asymmetries in the effects of campaigning or presidential approval. Do Democratic or Republican voters respond more strongly to either of these influences? Is campaigning's effect primarily one of mobilization—for example, Democratic campaigning encourages Democrats to vote—or one of demobilization—where Democratic campaigning discourages Republicans from voting?

To answer these questions, we rely on different dependent variables: the percentage of Democrats among voters and the percentage of Republicans among voters. These are computed based on the combined number of Democrats, Republicans, and Independents. Second, we disaggregate the campaign advantage measure into separate measures of Democratic and Republican campaign activity.¹² We then estimated two sets of models. The first uses the original campaign advantage measure, but disaggregates voters into Democrats and Republicans. The second disaggregates the campaign advantage measure as well. The models are identical in every other respect to Model 1 of Table 1.

Table 3 presents the results. Overall, these models reveal little evidence of asymmetry in the correlates of partisan turnout. In the first pair of models, the effect of campaign advantage is virtually identical in magnitude among both Democrats and Republicans. As Democratic campaigning increases relative to Republican campaigning, the Democratic share of the electorate increases ($b = 0.02$) and the Republican share of the electorate decreases ($b = -0.03$). The effects of presidential approval and state economic change, while statistically insignificant, are also symmetrical. In the second set of models, the separate measures of Democratic and Republican campaign activity also have nearly equal and opposite effects on the proportion of Democrats and Republicans. Again, the effects are not exactly equal, but the differences disappear after rounding. Voters in each party appear to respond in similar ways to the parties' respective campaign activities.

Discussion

Thus far we have shown that the partisan composition of the electorate is associated with the partisan balance of campaign activity. Two issues may complicate this story, however—the first concerning the apparent effect of campaign activity and the second concerning the interpretation of any shifts in partisan turnout. Because both issues are central to our empirical story, we dwell on them at some length.

The apparent effect of campaign activity could be an artifact of endogeneity: the balance of campaign activity in a particular state may anticipate, rather than affect, the partisan complexion of the electorate. This could occur in one of two ways. First, the balance of campaign spending might simply tilt in favor of the party that

¹² More specifically, for each party, we standardized each of the measures that comprise this index (spending on various races, presidential ads and visits, and party expenditures) and then combined those into an index.

Table 3 Models of partisan turnout, disaggregating partisan turnout and campaign activity

	(1)		(2)	
	Dem %	Rep %	Dem %	Rep %
Campaign advantage	0.02*	−0.03*		
	[0.01]	[0.01]		
Democratic campaign activity			0.01**	−0.01*
			[0.00]	[0.00]
Republican campaign activity			−0.01**	0.01*
			[0.00]	[0.00]
Presidential approval	0.03	−0.03	0.03	−0.04
	[0.03]	[0.03]	[0.03]	[0.03]
Economy	−0.07	−0.08	−0.07	−0.07
	[0.12]	[0.12]	[0.12]	[0.11]
President's party	−0.95	1.3	−1.14	1.49
	[1.41]	[1.15]	[1.42]	[1.22]
Constant	37.36**	35.63**	37.27**	35.70**
	[0.04]	[0.04]	[0.07]	[0.07]
Observations	269	269	269	269
Number of states	50	50	50	50
R-squared	0.06	0.09	0.08	0.07

Cell entries are least squares coefficients, with robust standard errors in parentheses. The dependent variables are the percentages of exit poll respondents who are Democratic or Republican in each state-year. These models also include fixed effects for states, as in Model 1 of Table 1. ⁺ $p < 0.10$; * $p < 0.05$; ** $p < 0.01$

dominates a state's electorate. Indeed, at the bivariate level, there are statistically significant correlations between the partisan balance of campaigning and the macro-partisanship of a particular state. However, the relationship between macro-partisanship and campaigning becomes statistically insignificant once the other attributes of states are controlled for via state-level fixed effects.¹³ This suggests that our fixed-effects models in Tables 1, 2, and 3 account for this species of endogeneity. Moreover, as shown in Table 1, our results are robust when state-level macropartisanship is included in the model.¹⁴

While this first kind of endogeneity is “spatial” in character, there are two sorts of “temporal” endogeneity as well. First, campaigns may steer their resources toward states whose electorates are changing in a favorable direction. For example, if the underlying partisanship of a state shifts toward the Democrats, Democratic candidates in that state may invest more resources. Second, both campaign activity and partisan turnout might respond to additional unmeasured variables that reflect the broader political climate in any given election year. We tested for the first

¹³ These results, and all others reported in this section, are also available from the authors. We do not present them here only for reasons of space.

¹⁴ The effects of campaign activity are also statistically significant when we measured partisan turnout as the deviation between the exit poll percentages and the underlying macro-partisanship of the state.

possibility by regressing the change in the state's balance of campaigning on the change in the underlying macro-partisanship of the state. Change was measured from one election to the next—i.e., we first-differenced both campaign activity and macro-partisanship. In this model, there was no statistically significant relationship, suggesting that campaign activity is not necessarily so sensitive to shifts in the underlying partisanship of states. Thus, this kind of endogeneity may not be prevalent.

To address the second form of temporal endogeneity, we added a series of dummy variables for election years to the models presented in Table 1. This approach assumes that any omitted variables that capture the broader political climate have a similar effect on all states. It also addresses the first form of temporal endogeneity to the extent that any shifts in underlying partisanship have been uniform across states. Again, the effect of campaign activity was unchanged. The same is true if we include *national* macro-partisanship in the model, which itself may proxy other changes in the political climate.

Finally, to address both spatial and temporal endogeneity, we leveraged the presence of divided Senate delegations. Although relatively few states have divided delegations, these states enable us to better isolate the effect of campaign activity. Adjacent elections that feature incumbent Senators of opposite parties allow us to hold many state-level attributes constant. Moreover, the schedule of the races themselves is exogenous, and while a Senator's decision to run for reelection is partly endogenous to factors such as any underlying shifts in the partisan complexion of the state, most of the difference in the balance of campaign activity between two such incumbents and their challengers likely reflects the advantages of incumbency. We also chose to compare Senate incumbents in sequential midterm elections, to avoid the confounding effects of presidential campaigning in the states. The most significant limitation of this approach is the small number of cases. Relatively few states in these years had divided senate delegations, and often there is no midterm exit poll in states without a Senate race.¹⁵ This left us with only seven races to examine. Nevertheless, the relationship between changes in Senate incumbency and changes in partisan turnout is in the direction we would expect: a change in Senate incumbency is associated with a shift of about six points in partisan turnout. In other words, if a Democratic Senator ran in 1998 after a Republican Senator had run in 1994, we would expect the partisan complexion of the electorate in 1998 to shift six points in the Democrats' favor, compared to 1994. We also expanded this analysis to include years when there was an open-seat Senate race or no Senate race in a particular state, assuming that partisan turnout in a year with, say, a Republican incumbent should tilt more toward the Republicans than it does in years with an open-seat race or no race. This brings the total number of races examined to 16, and again the shift in the partisan turnout is in the direction of the incumbent's party; the average shift is seven points. Of course, an analysis of so few

¹⁵ This analysis is somewhat in the spirit of Levitt (1994), who also sought to mitigate endogeneity by focusing on a small subset of races with fortuitous but relatively rare characteristics (in his analysis, races which featured the same opposing candidates more than once).

races is far from conclusive, but, in combination with our central analysis, it suggests that the effects of campaign activity on partisan turnout are real.

The second question concerns how we interpret the effect of campaign activity on partisan turnout. We have framed its effect as one of *mobilization*—that is, of getting friendly partisans to the polls. But could the effect be one of *conversion* rather than mobilization? In other words, could partisan turnout vary because people actually change their party identification? If the frequency of conversion were high enough, then any changes in partisan turnout could actually reflect changes in partisanship among an unchanging population of voters, in which case no mobilization would be occurring.

In some sense, either mobilization or conversion would represent a notable and important consequence of campaign activity. However, we suspect that mobilization is far more prevalent than conversion. First, the individual-level stability of party identification is more impressive than its malleability or endogeneity (Brody and Rothenberg 1988; Green et al. 2002; Jennings and Markus 1984; Markus 1982). But even if we take apparent shifts in party identification at face value, genuine defections from the Democratic Party to the Republican Party, or vice versa, are extremely rare. For example, in the 2000–2004 American National Election Study panel, only 1.4% of the sample shifted between the Democratic and Republican parties over this 4-year period. More common are shifts in the strength of partisanship (e.g., from strong Democrat to weak Democrat) or movement in and out of the “Independent” category. In the ANES panel, 2.5% of the sample consisted of partisans who moved into the independent category; 10% consisted of independents who moved into a partisan camp. We should also note that these raw percentages do not reflect corrections for measurement error, which would likely reduce the apparent instability of party identification (Green and Palmquist 1990).

This is particularly important to note because the exit poll measure of party identification does not capture strength of partisanship but merely allows respondents to categorize themselves as a Democrat, Republican, Independent, or something else. Thus, if shifts in partisan turnout were attributable to conversion, these shifts would most likely involve Independents’ moving into the party that is favored by fundamental conditions or that dominates the campaign. We tested this possibility by modeling the fraction of exit poll respondents who identify as Independent. If campaign activity leads some Independents to switch to a party, then the absolute value of the partisan balance in campaigning should be negatively associated with the fraction of Independents in the exit poll sample. That is, the more campaign activity favors one party over the other (regardless of which party it is), the fewer Independents there should be. We tested similar hypotheses involving both presidential approval and the state of the economy, again taking the absolute value of the relevant measures in our original models.¹⁶ A fixed-effects model of the

¹⁶ Recall that both measures were scaled such that negative values indicated conditions more favorable to the Republican Party and positive values indicated conditions more favorable to the Democratic Party. Thus, as the absolute value of these measures increases, one of the two parties is advantaged, which should lead to a decline in the number of Independents if some of them are switching parties.

fraction of Independents, including these three measures, produced no statistically significant relationships.

However, this finding may not imply the absence of conversion, as long as the number of independents moving to the favored party is roughly equal to the number of partisans from the unfavored party who become independents. To provide one final test, we examined the impact of campaign activity on state-level macro-partisanship, which we previously included as an independent variable (see Table 1). If campaign activity is converting voters rather than mobilizing them, its effects should also be manifest in overall macro-partisanship, not just the partisan complexion of voters. We estimated similar models of macro-partisanship, drawing on the specification in the first model of Table 1, but found no statistically significant relationship using either contemporaneous or lagged measures of campaign activity.¹⁷

Ultimately, while we cannot provide a definitive test of the prevalence of mobilization and conversion, our findings, combined with the accumulated evidence from research on party identification, suggests that variation in partisan turnout is due mostly to mobilization.

Conclusion

Does the partisan composition of the electorate respond to campaigns? Our findings suggest that it does. The more one party dominated the campaign, the greater the proportion of its supporters who went to the polls. Our findings thus confirm previous research that emphasizes the mobilizing effects of campaigning on partisans (Holbrook and McClurg 2005; Shaw 1999). To date this kind of partisan mobilization has received little systematic attention, perhaps because political scientists have often doubted that campaigns affect electoral behavior. Because we have extended previous analyses—both across time (1988–2006) and across political “space” (including both presidential and other races)—this mobilizing effect of campaigns rests on even firmer footing. Perhaps most important, we are able to show that campaigning matters over and above other political conditions, such as presidential approval and economic health.

In fact, the weak effects of these fundamental political conditions are a surprising result from our analysis. Neither presidential approval nor the state of the economy has a consistent, statistically significant effect on partisan turnout similar to the effect of campaign activity. We stop short of dismissing the effects of these variables, for two reasons. First, there is some evidence that the size of their effects varies across specifications. In particular, state-level economic conditions appear to

¹⁷ A possible rejoinder is that voters could be more attentive and sensitive to campaign stimuli, rendering them more likely than the population as a whole to shift their party identification. We think this is unlikely. Habitual voters are, on average, more attentive to politics, but they are also more opinionated. Indeed, further analysis from the 2000–2004 ANES panel suggests that respondents who reported voting in both 2000 and 2004 were even less likely than the entire sample to shift their party identification.

have a larger effect in midterms than in presidential years. Second, while our measures of the fundamentals are standard in the campaigns and elections literature, it is always possible that we have omitted a variable that would capture the true relationship between the fundamentals and partisan turnout. However, the robustness of our campaign activity measure across multiple specifications suggests that any measure of the fundamentals that future research employs is unlikely to account for the independent effect of campaigns.

An important implication is that things parties and candidates can control—the time, energy, and money they invest in speaking to voters and encouraging them to participate—matter to mobilization. These efforts do not simply capitalize on the fundamentals already in place (at least not the ones we account for), but have an apparent effect of their own. Moreover, their effects appear not only amidst the pomp and circumstance of presidential campaigns but during midterm elections as well. Our results also suggest that campaign resources can both mobilize fellow partisans and “de-mobilize” enemy partisans. Our data do not allow us to determine how much was spent on mobilization per se, and how exactly these monies were allocated to different sub-groups of voters. Nor do our data illuminate the individual-level psychological mechanisms that undergird these aggregate-level relationships. But the overall results suggest that campaign activities do alter the partisan complexion of voters. Simply put, if you want more of “your people” at the polls, you need to spend and do more than your opponent.

There is, however, an important caveat, if not an irony, to this finding. The ability of campaigns to mobilize their partisans depends on this asymmetry of resources and effort. But large asymmetries are only likely to exist in races that are fairly uncompetitive to begin with. In competitive races, the expenditures of the opposing sides will be closer to parity. It will be the hardest to obtain a competitive advantage through partisan mobilization precisely in those races where such an advantage would be most crucial to the overall outcome. But even small shifts in the composition of the electorate can matter at the margins, since the fair-weather fans behind them are probably reliable supporters of their party’s candidates when they choose to go to the polls. A useful extension of this analysis would be to estimate the impact of shifts in partisan complexion on election outcomes. Moreover, these shifts in the partisan complexion of voters might coincide with shifts in *ideological* complexion—i.e., a changing balance of moderates and ideologues within each party’s voters. Shifts of this kind cannot be measured with the data we have, but future research might identify such patterns.

The movement in partisan turnout that we have found here suggests an important role for short-term turnout effects in contemporary American politics. As many have long suspected, elections are about more than just convincing voters. They are also about mobilizing the voters who have already been convinced.

Appendix

See Table 4.

Table 4 Distribution of states in exit poll data

State	1988	1990	1992	1994	1996	1998	2000	2002	2004	2006	Total
Alabama	0	1	1	0	1	1	1	0	1	0	6
Alaska	0	0	1	0	1	0	1	0	1	0	4
Arizona	0	0	1	1	1	1	1	0	1	1	7
Arkansas	0	0	1	0	1	1	1	0	1	0	5
California	1	0	1	1	1	1	1	1	1	1	9
Colorado	0	1	1	0	1	1	1	0	1	0	6
Connecticut	1	0	1	1	1	1	1	0	1	1	8
Delaware	0	1	1	1	1	0	1	0	1	0	6
Florida	1	0	1	1	1	1	1	1	1	1	9
Georgia	0	0	1	0	1	1	1	0	1	1	6
Hawaii	0	1	1	0	1	0	1	0	1	1	6
Idaho	0	1	1	0	1	1	1	0	1	0	6
Illinois	1	1	1	0	1	1	1	1	1	1	9
Indiana	1	1	1	0	1	1	1	0	1	0	7
Iowa	1	1	1	0	1	1	1	0	1	0	7
Kansas	0	1	1	0	1	1	1	0	1	0	6
Kentucky	0	1	1	0	1	1	1	1	1	0	7
Louisiana	0	0	1	0	1	1	1	0	1	0	5
Maine	0	1	1	1	1	1	1	0	1	1	8
Maryland	1	0	1	1	1	1	1	0	1	1	8
Massachusetts	1	1	1	1	1	1	1	0	1	1	9
Michigan	1	1	1	1	1	1	1	0	1	1	9
Minnesota	1	1	1	1	1	1	1	0	1	1	9
Mississippi	1	0	1	0	1	0	1	0	1	0	5
Missouri	1	0	1	1	1	1	1	0	1	1	8
Montana	0	1	1	1	1	0	1	0	1	1	7
Nebraska	0	1	1	1	1	1	1	0	1	1	8
Nevada	1	0	1	1	1	1	1	0	1	1	8
New Hampshire	0	1	1	0	1	1	1	0	1	0	6
New Jersey	1	1	1	1	1	0	1	0	1	1	8
New Mexico	1	1	1	1	1	1	1	0	1	1	9
New York	1	0	1	1	1	1	1	0	1	1	9
North Carolina	1	1	1	0	1	1	1	0	1	0	7
North Dakota	0	0	1	0	1	1	1	0	1	1	6
Ohio	1	0	1	1	1	1	1	1	1	1	9
Oklahoma	0	1	1	1	1	1	1	0	1	0	7
Oregon	1	1	1	0	1	1	1	0	1	0	7
Pennsylvania	1	0	1	1	1	1	1	1	1	1	9
Rhode Island	0	1	1	0	1	1	1	0	1	1	7
South Carolina	0	1	1	0	1	1	1	0	1	0	6
South Dakota	0	1	1	0	1	0	1	0	1	0	5

Table 4 continued

State	1988	1990	1992	1994	1996	1998	2000	2002	2004	2006	Total
Tennessee	0	1	1	1	1	1	1	0	1	1	8
Texas	1	1	1	1	1	1	1	1	1	1	10
Utah	0	0	1	0	1	1	1	0	1	1	6
Vermont	1	0	1	0	1	1	1	0	1	1	7
Virginia	0	0	1	1	1	0	1	0	1	1	6
Washington	1	0	1	1	1	1	1	0	1	1	8
West Virginia	0	1	1	0	1	0	1	0	1	1	6
Wisconsin	1	0	1	1	1	1	1	0	1	1	8
Wyoming	0	1	1	1	1	1	1	0	1	1	8
Total	24	29	50	26	50	41	50	8	50	32	360

Note: 1 indicates a state-level sample in that year

References

- Ansolabehere, S., & Iyengar, S. (1995). *Going negative: How attack ads shrink and polarize the electorate*. New York: Free Press.
- Arclusus, F., & Meltzer, A. H. (1975). The effect of aggregate economic variables on congressional elections. *The American Political Science Review*, 69(4), 1232–1239.
- Beck, N., & Katz, J. N. (1995). What to do (and not to do) with time series cross-section data. *The American Political Science Review*, 89, 634–647.
- Beyle, T. & Jensen, J. M. (2007). Gubernatorial campaign expenditures database. Cooperative project of the University at Albany, SUNY, and the University of North Carolina at Chapel Hill.
- Brody, R. A., & Rothenberg, L. S. (1988). The instability of partisanship: An analysis of the 1980 presidential election. *British Journal of Political Science*, 18(4), 445–465.
- Caldeira, G. A., & Patterson, S. C. (1982). Contextual influences on participation in US state legislative elections. *Legislative Studies Quarterly*, 7(3), 359–381.
- Campbell, A. (1960). Surge and decline: A study of electoral change. *Public Opinion Quarterly*, 24, 397–418.
- Campbell, J. E. (1997). *The presidential pulse of congressional elections*. Lexington: University of Kentucky.
- Coates, D., & Humphreys, B. R. (2005). Novelty effects of new facilities on attendance at professional sporting events. *Contemporary Economic Policy*, 23(3), 436–455.
- Converse, P. (1964). The nature of belief systems in mass publics. In D. E. Apter (Ed.), *Ideology and discontent*. New York: Free Press.
- Cox, G. W., & Munger, M. C. (1989). Closeness, expenditures, and turnout in the 1982 US house elections. *The American Political Science Review*, 83(1), 217–231.
- Erikson, R. S., Wright, G. C., & McIver, J. P. (1989). Political parties, public opinion, and state policy in the United States. *The American Political Science Review*, 83(3), 729–750.
- Gilliam, F. D. (1985). Influences on voter turnout for US house elections in non-presidential years. *Legislative Studies Quarterly*, 10(3), 339–351.
- Gimpel, J. G., Kaufmann, K. M., & Pearson-Merkowitz, S. (2007). Battleground states versus blackout states: The behavioral implications of modern presidential campaigns. *The Journal of Politics*, 69(3), 786–797.
- Green, D. P., & Gerber, A. S. (2004). *Get out the vote!: How to increase voter turnout*. Washington, DC: Brookings Institution Press.
- Green, D. P., & Palmquist, B. (1990). Of artifacts and partisan instability. *American Journal of Political Science*, 34(3), 872–902.
- Green, D. P., Palmquist, B., & Schickler, E. (2002). *Partisan hearts and minds: Political parties and the social identities of voters*. New Haven: Yale University Press.

- Holbrook, T. M., & McClurg, S. D. (2005). The mobilization of core supporters: Campaigns, turnout, and electoral composition in United States presidential elections. *American Journal of Political Science*, 49(4), 689–703.
- Jackson, R. A. (2002). Gubernatorial and senatorial campaign mobilization of voters. *Political Research Quarterly*, 55(4), 825–844.
- Jacobson, G. C., & Kernell, S. (1983). *Strategy and choice in congressional elections*. New Haven: Yale University Press.
- Jennings, M. K., & Markus, G. B. (1984). Partisan orientations over the long haul: Results from the three-wave political socialization panel study. *The American Political Science Review*, 78(4), 1000–1018.
- Levitt, S. D. (1994). Using repeat challengers to estimate the effect of campaign spending on election outcomes in the US house. *Journal of Political Economy*, 102(4), 777–798.
- Lewis-Beck, M. S., & Tom Rice, W. (1992). *Forecasting elections*. Washington, DC: Congressional Quarterly Press.
- Marcus, G. E., Russell Neuman, W., & MacKuen, M. (2000). *Affective intelligence and political judgment*. Chicago: University of Chicago Press.
- Markus, G. B. (1982). Political attitudes during an election year: A report on the 1980 NES Panel study. *The American Political Science Review*, 76, 538–560.
- Paul, H. J. (2002). The impact of campaign appearances in the 1996 election. *Journal of Politics*, 64(3), 904–913.
- Richard N., Beyle T., & Sigelman L. (2007). U.S. Officials' Job Approval Ratings. Cooperative project of the University of Rochester, The University of North Carolina at Chapel Hill, and George Washington University.
- Rosenstone, S. J., & Hansen, J. M. (1993). *Mobilization, participation, and democracy in America*. New York: Macmillan.
- Shaw, D. R. (1999). The effect of TV Ads and candidate appearances on statewide presidential votes, 1988–96. *The American Political Science Review*, 93(2), 345–361.
- Shaw, D. R. (2006). *The race to 270: The Electoral College and the campaign strategies of 2000 and 2004*. Chicago: Chicago University Press.
- Sigelman, L., & Jewell, M. E. (1986). From core to periphery: A note on the imagery of concentric electorates. *The Journal of Politics*, 48(2), 440–449.
- Verba, S., Scholzman, K. L., & Brady, H. E. (1995). *Voice and equality: Civic voluntarism in American politics*. Cambridge, MA: Harvard University Press.
- Wilson, S. E., & Butler, D. M. (2007). A lot more to do: The sensitivity of time-series cross-section analysis to simple alternative specifications. *Political Analysis*, 15(2), 101–123.
- Wolfinger, R. E., & Rosenstone, S. J. (1980). *Who Votes?*. New Haven: Yale University Press.